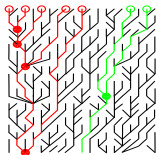


# Evolution of the rate of evolution

—

## An analytical solution to the compound Poisson process



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# Outline

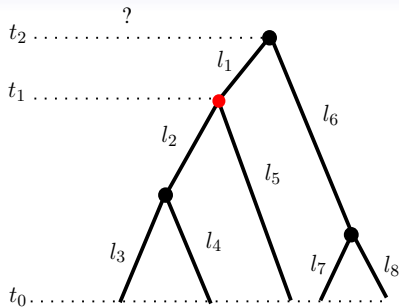
Models of evolution of the rate of evolution

The compound Poisson process: an analytical solution

## A bit of history...

- Linus Pauling and Emile Zuckerkandl (1962): “*molecular clock hypothesis*”.
- Allan Wilson (1967): molecular dating under the molecular clock assumption.
- 30 years passed...
- Michael Sanderson (1997) and Jeffrey Thorne (1998): estimation of evolutionary divergence times without the restriction of a uniform rate across lineages.

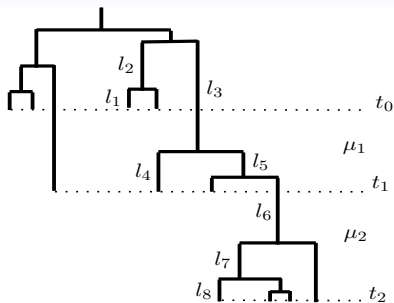
# Molecular clock rate and time estimation



$$l_5 = \mu \times (t_1 - t_0)$$

$$\hookrightarrow \mu = \frac{l_5}{t_1 - t_0}$$

$$t_2 = \frac{l_1 + l_2 + l_3}{\mu} + t_0$$



$$\mu_1 = \frac{l_4 + l_3 - l_1 - l_2}{t_0 - t_1}$$

$$\mu_2 = \frac{l_5 + l_6 + l_7 + l_8 - l_4}{t_1 - t_2}$$

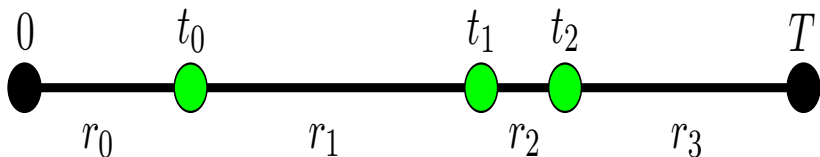
# Beyond the molecular clock

- *Local clocks*
  - Substitution rate is organised into a small number of classes,
  - Assign each branch to one of these classes.
- *Penalized likelihood*
  - $\Psi(R, T)$ : penalty term for rate changes,
  - Maximise  $\log(P(D|R, T)) - \lambda\Psi(R, T)$ .
- *Bayesian approaches*
  - Explicit stochastic models of the evolution of the substitution rate.
  - Rate trajectory is continuous or discrete.

# Models of rate evolution (1/2)

- Log-normal model
  - $\mu$  is the mean of the rate at the nodes that begin and end the branch ( $r(0)$  and  $r(T)$ ).
  - $\log(r(T)) \sim \mathcal{N}(\log(r(0)), \nu T)$ .
  - Logarithm of the rate undergoes *Brownian motion*.
  - Correlation of mean rates on adjacent branches.
- Exponential model
  - $\mu \sim \text{Exp}(\phi)$ .
  - No correlation of mean rates.
  - Shape of the distribution does not depend on time duration.

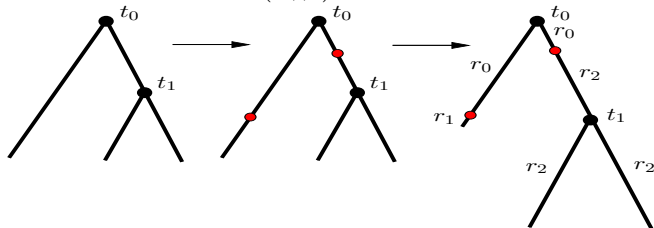
## Models of rate evolution (2/2)



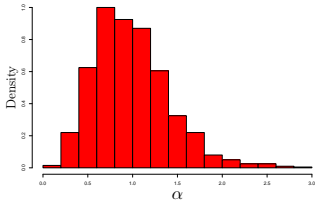
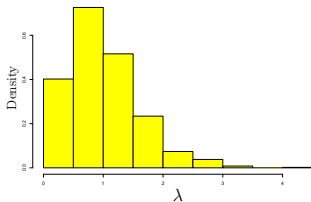
- Compound Poisson process
  - Rates change in discrete jumps.
  - $r(t) \sim \Gamma(\alpha, \beta)$
  - Number of jumps:  $n(T) \sim \text{Poisson}(\lambda T)$
  - Correlation of mean rates across branches: governed by  $\lambda$ .
  - $\lambda T$  large: distribution of mean rate is approximately Normal.

# Implementation of the compound Poisson process

- “Jump” event:  $Poisson(\lambda\Delta t)$
- Substitution rates:  $\Gamma(\alpha, \beta)$



- MCMC  $\rightarrow$  posterior distribution of  $\lambda$  and  $\alpha$





# Advantages and drawbacks

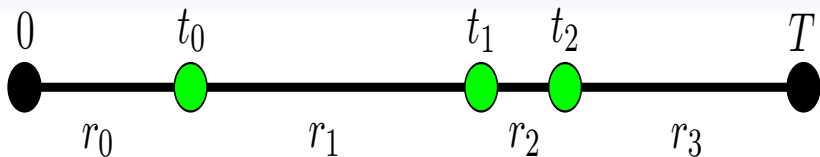
- Log-normal
  - Computationally tractable
  - Crude (deterministic) description of the mean rates.
  - Biologically relevant ?
- Exponential
  - Computationally tractable.
  - Distribution of mean substitution rate does not depend on time duration.
  - No correlation of mean rates across branches.
- Compound Poisson
  - Description of rate changes plausible from a biological perspective.
  - Elegant way to account for correlation of mean rates across branches.
  - No analytical solution.

# Outline

Models of evolution of the rate of evolution

The compound Poisson process: an analytical solution

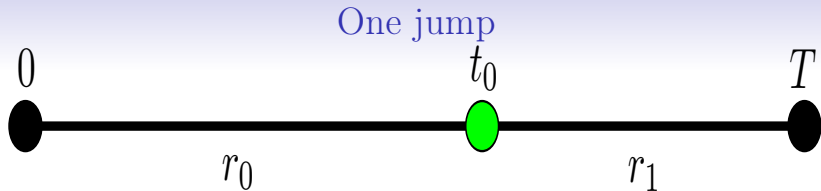
## First question



- $r_i \sim \Gamma(\alpha, \beta)$ . Hence,  $E(r_i) = \alpha\beta$ ,  $V(r_i) = \alpha\beta^2$ .
- $n \sim \text{Poisson}(\lambda T)$ .
- $\mu = \sum_{i=0}^n k_i r_i$ , where  $k_i = \frac{\Delta t_i}{T}$ .

What is the distribution of  $\mu$  ?

- Work out the distribution of  $\mu$  for a given value of  $n$ .
- $\mu = \sum_{i=0}^n k_i r_i$  is well approximated by a Gamma distribution.

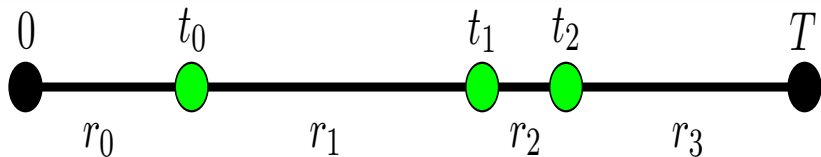


- $\mu = k_0 r_0 + (1 - k_0) r_1$
- Distribution of  $t_0 = k_0 T$  ?

$$\begin{aligned}
 P(t_0 = x | n = 1) &= \frac{\lambda e^{-\lambda x} \times e^{-\lambda(T-x)}}{\lambda T e^{-\lambda T}} \\
 &= \frac{1}{T}.
 \end{aligned}$$

- $k_0 \sim U[0, 1] \rightarrow E(k_0) = \frac{1}{2}$  and  $V(k_0) = \frac{1}{12}$ .
- $E(\mu) = E(k_0)E(r_0) + E(1 - k_0)E(r_1) = \alpha\beta$ .
- $V(\mu) = V(k_0 r_0) + V((1 - k_0) r_1) + 2Cov(k_0 r_0, (1 - k_0) r_1) = \frac{2}{3}\alpha\beta^2$ .

$n \geq 1$  jumps



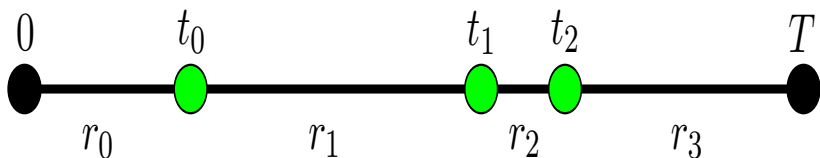
- Distribution of  $k_0$  ?

$$\begin{aligned} P(t_0 = x | n = y) &= \frac{\lambda e^{-\lambda x} \times (\lambda(T-x))^{y-1} e^{-\lambda(T-x)}}{(\lambda T)^y e^{-\lambda T} / y!} \\ &= \frac{y}{T^y} (T-x)^{y-1}. \end{aligned}$$

- After little algebra...

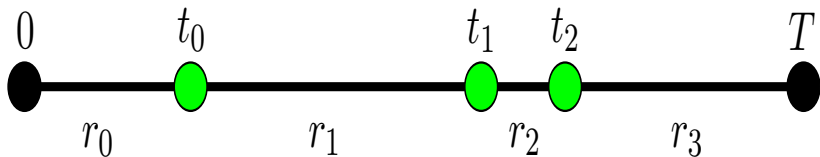
- $E(k_0) = \frac{1}{n+1}$ ,
- $E(k_0^2) = \frac{2}{(n+1)(n+2)}$ .

$n \geq 1$  jumps



- $\mu = k_0 r_0 + k_1 r_1 + k_2 r_2 + k_3 r_3.$
- $\mu_n = k_0 r_0 + (1 - k_0) \mu_{n-1}.$
- $E(\mu_n) = E(k_0)E(r_0) + E(1 - k_0)E(\mu_{n-1}) \rightarrow \boxed{E(\mu_n) = \alpha\beta}.$

$n \geq 1$  jumps



- The variance is a bit more challenging but can be done.

$$V(\mu_n) = \frac{2\alpha\beta^2 + n(n+1)V(\mu_{n-1})}{(n+1)(n+2)}$$

- Solve the recursion:

$$\boxed{V(\mu_n) = \frac{2}{n+2}\alpha\beta^2}$$

# Likelihood calculation

- Data:
  - $l$ , an expected number of substitutions.
  - $T$ , elapsed time.
- $\mu = l/T$
- Likelihood:

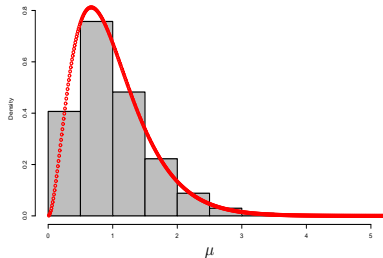
$$p_{\mu}(u|\lambda, \alpha, \beta, T) = \sum_{n=0}^{\infty} P(n|\lambda, T) p_{\mu_n}(u|\alpha, \beta, n)$$

- $P(n|\lambda, T)$ : Poisson distribution with mean and variance  $\lambda T$ .
- $p_{\mu_n}(u|\alpha, \beta, n)$ : Gamma distribution with mean  $\alpha\beta$ , and variance  $\frac{2}{n+2}\alpha\beta^2$ .

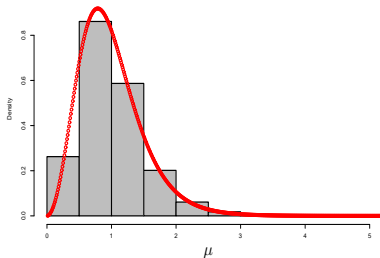


# The approximation seems good

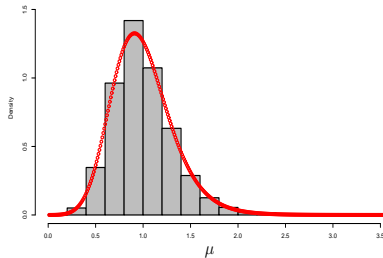
$\lambda = 1E-04$  ( $E(n) = 0.001$ )



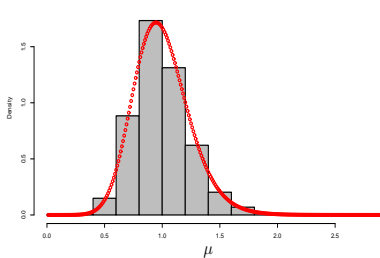
$\lambda = 0.1$  ( $E(n) = 1$ )



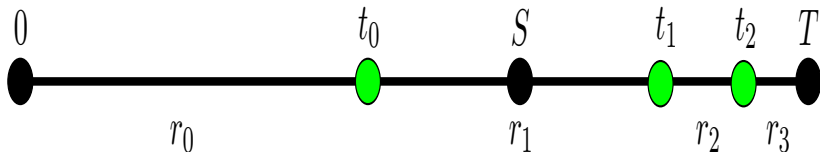
$\lambda = 0.5$  ( $E(n) = 5$ )



$\lambda = 1$  ( $E(n) = 10$ )



## Second question

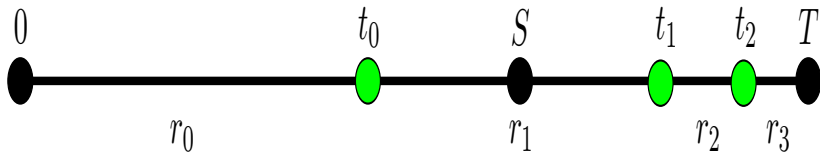


- Two adjacent time intervals:  $[0, S]$  and  $[S, T]$ .
- $\mu_1$  and  $\mu_2$  mean rates in  $[0, S]$  and  $[S, T]$  respectively.
- $\mu_1$  and  $\mu_2$  are correlated because of  $r_1$ .

What is the joint distribution of  $\mu_1$  and  $\mu_2$  ?

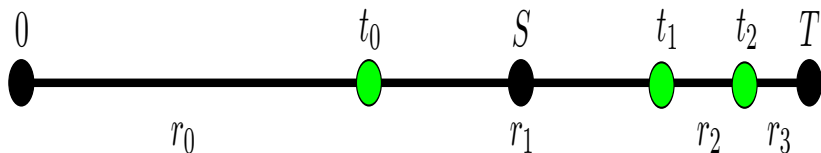
- Work out the density  $p_{\mu_2|\mu_1}(u_2|u_1, \lambda, \alpha, T - S)$ .

## Second question



- I was unable to find an analytical expression...
- First idea: integrate over  $t_0$  in  $[0, S]$ ,  $t_1$  in  $[S, T]$  and  $r_1$  in  $[0, \infty]$ ...
- ...didn't work.
- Second idea: use an approximation.
  - 'Many' jumps in  $[0, T]$ :  $\mu_1$  and  $\mu_2$  are independent.
  - No jump in  $[0, T]$ :  $p_{\mu_2|\mu_1}(u_2|u_1) = 1$  if  $u_2 = u_1$ .

## Second question

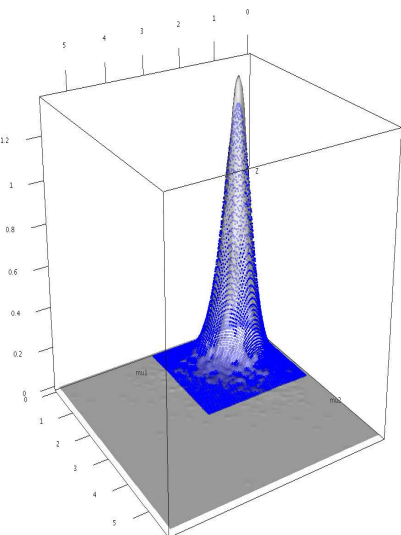
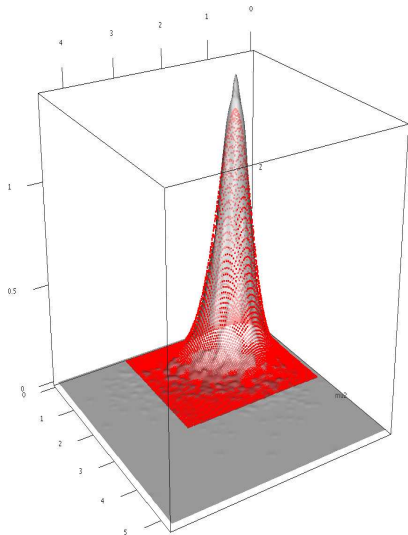


- Use a *mixture model*:
  - $\mu_2|\mu_1 \sim \mathcal{N}(\mu_1, 0.01)$  with probability  $P(n = 0|\lambda, T)$ ,
  - $\mu_2|\mu_1 \sim \mathcal{N}(\mu_1, 0.04)$  with probability  $P(n = 1|\lambda, T)$ ,
  - $\mu_2|\mu_1 \sim \mathcal{N}(\mu_1, 0.09)$  with probability  $P(n = 2|\lambda, T)$ ,
  - $\mu_2$  independent from  $\mu_1$  with probability  $P(n > 2|\lambda, T)$ .

$$p_{\mu_1, \mu_2}(u_1, u_2 | \lambda, \alpha, T), E(n) = 10$$

Mixture

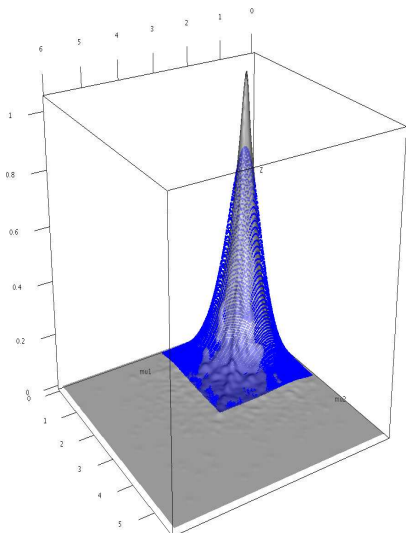
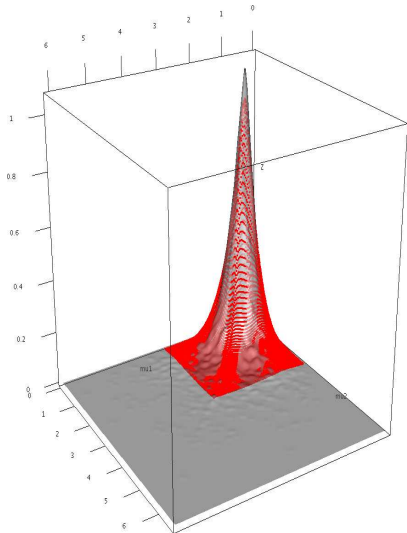
Independent



$$p_{\mu_1, \mu_2}(u_1, u_2 | \lambda, \alpha, T), E(n) = 4$$

Mixture

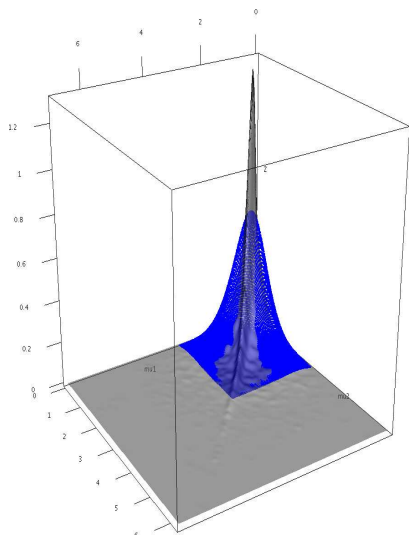
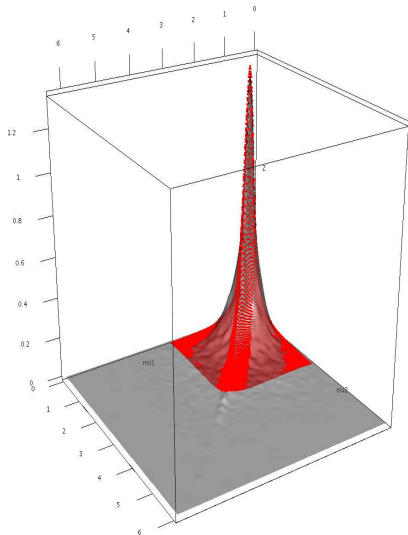
Independent



$$p_{\mu_1, \mu_2}(u_1, u_2 | \lambda, \alpha, T), E(n) = 2$$

Mixture

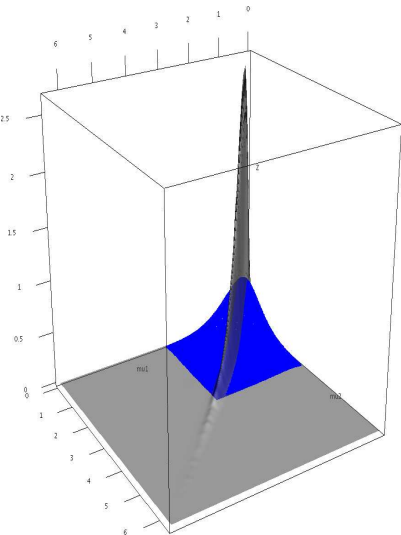
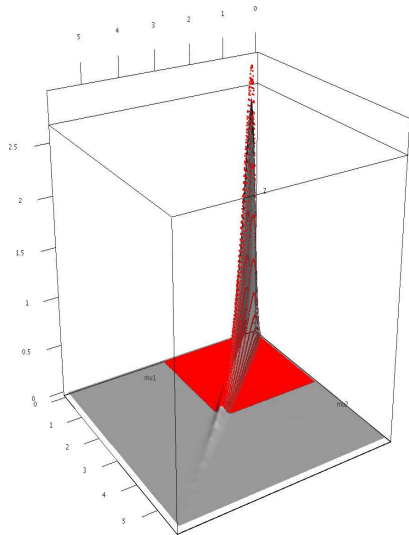
Independent



$$p_{\mu_1, \mu_2}(u_1, u_2 | \lambda, \alpha, T), E(n) = 0.002$$

Mixture

Independent





# Acknowledgements

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- Dumont d'Urville programme:
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